Fitting Multidimensional Latent Variable Models using an Efficient Laplace Approximation

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- Item Response Theory (IRT) plays nowadays a central role in the analysis and study of tests and item scores
- Application of IRT models can be found in many fields
 - ▷ psychometrics
 - > educational sciences
 - \triangleright sociometrics
 - \triangleright medicine
 - ▷...



- Standard IRT models are available in special-purpose software, such as BILOG & MULTILOG and in R
- For R more information can be found at: http://cran.r-project.org/web/views/Psychometrics.html



- A fundamental assumption behind these standard IRT models is *unidimensionality*:
 - b the interdependencies between the responses of each sample unit are explained by a *single* latent variable
- In some cases tests are designed to measure a single trait, e.g.,
 - ▷ reading ability
 - > environmental attitude
 - ▷...



- However, in many cases unidimensionality is too strict to be true, e.g.,
 - b tests measure different latent traits
 - * mathematics test: algebra, calculus, etc.
 - * types of depression: major depressive disorder, dysthymia, manic depression
 - b hierarchical/multilevel designs
 - * subjects are nested within clusters
 - * items are nested within different dimensions



- If there is a predominant general factor in the data, and dimensions beyond that major dimension are relatively small, then multidimensionality has a little effect on derived inferences
- However, if the unidimensionality assumption is seriously violated, then

▷ item parameter estimates will be biased, and

▷ the standard errors associated with ability parameter estimates will be too small



- Programme for International Student Assessment (PISA)
 - Iaunched by the Organization for Economic Co-operation and Development
 - collect data on student and institutional factors that can explain differences in student performance
 - b in 2003, 41 countries participated and the survey covered mathematics, reading, science, and problem solving
- Data features
 - > different dimensions: ability in mathematics, reading, science, problem solving
 - > hierarchical design: students nested in schools, schools nested in countries



- Aim: estimate item and ability parameters, taking into account covariates and the hierarchical design
- Using a multilevel analysis we will be able to simultaneously estimate the item and ability



- Problem: as we will illustrate fitting complex latent variable models is a computationally challenging task requiring a lot of computing time
- Our Aim: develop a computationally flexible approach that can fit latent variable models with complex latent structures in reasonable computing time
- Work in progress... (no results yet available)
 - promising results from the relevant framework of joint models for longitudinal and time-to-event data (with high-dimensional random effects)



- Notation:
 - $\triangleright \boldsymbol{y}_i$: vector of responses for the *i*th subject
 - $\triangleright \boldsymbol{z}_i$: vector of latent variables
- Basic assumption: conditional independence (CI)
 - \triangleright given the latent structure, we assume that the responses of the $i{\rm th}$ subject are independent

$$p(\boldsymbol{y}_i \mid \boldsymbol{z}_i) = \prod_{k=1}^p p(y_{ik} \mid \boldsymbol{z}_i)$$

where $p(\cdot)$ denotes a pdf



- In order CI to hold, a complex latent structure may be required
- A general definition of an IRT model

$$g\{E(\boldsymbol{y}_i \mid \boldsymbol{z}_i)\} = \boldsymbol{X}_i \boldsymbol{\beta}^{(x)} + \boldsymbol{Z}_i \boldsymbol{\beta}^{(z)}$$

where

 $\triangleright g(\cdot)$: link function

- $\triangleright \boldsymbol{X}_i$: design matrix for covariates
- $\triangleright \boldsymbol{Z}_i$: vector of latent variables
- $\triangleright \beta^{(x)}$: regression coefficients for covariates
- $\triangleright \beta^{(z)}$: regression coefficients for latent variables



• Examples:

 \triangleright dichotomous data – 1 level (*i* subject, *k* item)

$$\mathsf{logit}\{\mathsf{Pr}(y_{ik}=1 \mid \boldsymbol{z}_i, \boldsymbol{\theta})\} = \beta_0 + \beta_1 z_{i1} + \beta_2 z_{i2} + \ldots + \beta_q z_{iq}$$

q-latent-variable model



• Examples:

 \triangleright polytomous data (c = 1, 2, ...) – 2 levels (*i* subject in group *j*, *k* item)

$$\Pr(y_{ijk} = c \mid \boldsymbol{z}_i, \boldsymbol{\theta}) = \exp((a_k z_{ij} - b_{k,c-1})) - \exp((a_k z_{ij} - b_{k,c}))$$

Level I:
$$z_{ij} = \beta_{0j} + \beta_{1j}x_{1ij} + e_{ij}$$

Level II:
$$\beta_{0j} = \gamma_{00} + \gamma_{01}w_{1j} + u_{0j}$$

 $\beta_{1j} = \gamma_{10} + \gamma_{11}w_{1j} + u_{1j}$

 $e_{ij}, \ \boldsymbol{u}_i$ denote Error Terms



• Estimation of multidimensional IRT model is typically based on marginal maximum likelihood

$$\ell(\boldsymbol{\theta}) = \sum_{i=1}^{n} \log \int p(\boldsymbol{y}_i \mid \boldsymbol{z}_i; \boldsymbol{\theta}) \ p(\boldsymbol{z}_i; \boldsymbol{\theta}) \ d\boldsymbol{z}_i$$

where

 $\triangleright \boldsymbol{ heta}$ denotes the parameter vector

 $\triangleright p(\boldsymbol{y}_i \mid \boldsymbol{z}_i; \boldsymbol{\theta})$ denoted the density of the multidimensional IRT as introduced above \triangleright we assume that \boldsymbol{z}_i are distributed according to a parametric distribution \triangleright we integrate \boldsymbol{z}_i to obtain the marginal distribution for the observed responses



• Due to the fact that the integral

$$\int p(oldsymbol{y}_i | oldsymbol{z}_i) \; p(oldsymbol{z}_i) \; doldsymbol{z}_i$$

does not have a closed form solution

- Maximization of $\ell(\theta)$ is a computationally challenging task it requires a combination of
 - \triangleright numerical integration, and
 - ▷ numerical optimization



- For numerical optimization standard choices are
 - \triangleright EM algorithm (we treat z_i as 'missing values')
 - > Newton-type algorithms, such as Newton-Raphson or quasi-Newton
- Hybrid approaches that start with EM (as a refinement of the starting values) for a fixed number of iterations, and continue with quasi-Newton have also been successfully used



- For numerical integration standard choices are
 - \triangleright Monte Carlo
 - > (adaptive) Gauss-Hermite quadrature rule
- <u>However</u>, these are prohibitive when a moderate to high number of latent variables is considered















• An alternative solution instead of numerical integration is the Laplace approximation

$$p(\boldsymbol{y}_i; \boldsymbol{\theta}) = \int \exp\{\log p(\boldsymbol{y}_i \mid \boldsymbol{z}_i; \boldsymbol{\theta}) + \log p(\boldsymbol{z}_i; \boldsymbol{\theta})\} d\boldsymbol{z}_i$$
$$= \left[(2\pi)^{q/2} \det(\boldsymbol{\Sigma})^{-1/2} \exp\{\log p(\boldsymbol{y}_i \mid \hat{\boldsymbol{z}}_i; \boldsymbol{\theta}) + \log p(\hat{\boldsymbol{z}}_i; \boldsymbol{\theta})\} \right] \left(1 + O(p_i^{-1}) \right),$$

where

$$\triangleright \, \widehat{\boldsymbol{z}}_i = \operatorname*{argmax}_i \{ \log p(\boldsymbol{y}_i \mid \boldsymbol{z}_i) + \log p(\boldsymbol{z}_i) \}$$
$$\triangleright \, \boldsymbol{\Sigma} = -\nabla^2 \left\{ \log p(\boldsymbol{y}_i \mid \boldsymbol{z}_i) + \log p(\boldsymbol{z}_i) \right\} \Big|_{\boldsymbol{z}_i = \widehat{\boldsymbol{z}}_i}$$



- It requires a large number of repeated measurements per individual in order to provide a good approximation to the integral
- Contrary to Monte Carlo and Gaussian quadrature, in the Laplace approximation we cannot control the approximation error
- Therefore, it would be desirable to improve the approximation, especially for small to moderate number of repeated measurements per individual



• The score vector in latent variable models can be written in the form (Rizopoulos et al., JRSSB, 2009)

$$S_{i}(\boldsymbol{\theta}) = \sum_{i} \frac{\partial}{\partial \boldsymbol{\theta}} \log \int p(\boldsymbol{y}_{i} \mid \boldsymbol{z}_{i}; \boldsymbol{\theta}) p(\boldsymbol{z}_{i}; \boldsymbol{\theta}) d\boldsymbol{z}_{i}$$
$$= \sum_{i} \int \frac{\partial}{\partial \boldsymbol{\theta}} \left\{ \log p(\boldsymbol{y}_{i} \mid \boldsymbol{z}_{i}; \boldsymbol{\theta}) + \log p(\boldsymbol{z}_{i}; \boldsymbol{\theta}) \right\} p(\boldsymbol{z}_{i} \mid \boldsymbol{y}_{i}; \boldsymbol{\theta}) d\boldsymbol{z}_{i}$$

• Observed data score vector = expected value of complete data score vector *wrt* the posterior of the latent variables given the observed data



- Why is this useful
 - \triangleright easy to combine EM with quasi-Newton
 - ▷ enables a more efficient Laplace approximation



- EM algorithm for latent variable models
 - \triangleright maximize the expected value of the complete data log-likelihood (expectation is taken *wrt* the posterior of the latent variables given the observed data)

$$Q_i(\boldsymbol{\theta} \mid \boldsymbol{\theta}^*) = \int \log\{p(\boldsymbol{y}_i \mid \boldsymbol{z}_i; \boldsymbol{\theta}) p(\boldsymbol{z}_i; \boldsymbol{\theta})\} \ p(\boldsymbol{z}_i \mid \boldsymbol{y}_i; \boldsymbol{\theta}^*) \ d\boldsymbol{z}_i$$

 \bullet To maximize $Q(\cdot)$ we need to solve

$$\int \frac{\partial}{\partial \boldsymbol{\theta}} \Big\{ \log p(\boldsymbol{y}_i \mid \boldsymbol{z}_i; \boldsymbol{\theta}) + \log p(\boldsymbol{z}_i; \boldsymbol{\theta}) \Big\} p(\boldsymbol{z}_i \mid \boldsymbol{y}_i; \boldsymbol{\theta}^*) \, d\boldsymbol{z}_i = 0$$

which is $S_i(\boldsymbol{\theta})$



- Direct maximization for latent variable models using quasi-Newton
 - \triangleright maximize the observed data log-likelihood \Rightarrow solve the score equations $S_i(\theta) = 0$
- Therefore, both EM and quasi-Newton require calculation of the same function $S_i(\boldsymbol{\theta})$
 - b take into advantage of the stability of EM during the first iteration, and later change to quasi-Newton which has better convergence rate



• Fitting latent variable models under MML requires calculations of the form

$$\int A(\boldsymbol{z}_i) \ p(\boldsymbol{z}_i \mid \boldsymbol{y}_i) \ d\boldsymbol{z}_i,$$

where $A(\boldsymbol{z}_i) = \partial \{\log p(\boldsymbol{y}_i \mid \boldsymbol{z}_i; \boldsymbol{\theta}) + \log p(\boldsymbol{z}_i; \boldsymbol{\theta})\} / \partial \boldsymbol{\theta}$

• Note that the above can be written as

$$E\{A(\boldsymbol{z}_i)\} = \frac{\int A(\boldsymbol{z}_i) \ p(\boldsymbol{y}_i \mid \boldsymbol{z}_i) \ p(\boldsymbol{z}_i) \ d\boldsymbol{z}_i}{\int p(\boldsymbol{y}_i \mid \boldsymbol{z}_i) \ p(\boldsymbol{z}_i) \ d\boldsymbol{z}_i}$$



- If we apply the standard Laplace approximation in the numerator and denominator of $E\{A(\boldsymbol{z}_i)\}$, then the $O(p_i^{-1})$ terms cancel out, which leads to a $O(p_i^{-2})$ approximation
- This approximation has been used for Bayesian computations (Tierney et al., JASA, 1989)
- <u>Caveat:</u> it can only be applied for positive functions
 - \triangleright however, $A(\pmb{z}_i)$, which is the complete data score vector, is not restricted to be positive



• Write the previous equation as

$$E\left\{A(\boldsymbol{z}_{i})\right\} = \frac{d}{ds}\log E\left[\exp\left\{sA(\boldsymbol{z}_{i})\right\}\right]\Big|_{s=0}$$

• Then we obtain the approximation

$$E\left\{A(\boldsymbol{z}_i)\right\} = \left\{A(\widehat{\boldsymbol{z}}_i) + \frac{d}{ds}\log\det(\boldsymbol{\Sigma}_s)^{-1/2}\Big|_{s=0}\right\} \left(1 + O(p_i^{-2})\right),$$

where

$$\triangleright \boldsymbol{\Sigma}_{s} = -\nabla^{2} \left\{ sA(\boldsymbol{z}_{i}) + \log p(\boldsymbol{y}_{i} \mid \boldsymbol{z}_{i}) + \log p(\boldsymbol{z}_{i}) \right\} \Big|_{\boldsymbol{z}_{i} = \widehat{\boldsymbol{z}}_{i}^{(s)}}$$

 $\triangleright \widehat{\boldsymbol{z}}_i$ same as in the simple Laplace approximation



- The enhanced Laplace approximation is
 - ▷ the simple Laplace approximation,

 \triangleright and differentiation of $\{\log \det(\mathbf{\Sigma}_s)^{-1/2}\}$ wrt~s

$$\frac{\partial}{\partial s_k} \log \det(\Sigma_s)^{-1/2} = -\frac{1}{2} \operatorname{tr} \left(\left. \Sigma^{-1} \frac{\partial}{\partial s_k} \Sigma_s \right|_{s=0, z_i = \hat{z}_i} \right)$$

• Features:

- \triangleright it is rather technical (you can get lost in the derivatives of $\{\log \det(\Sigma_s)^{-1/2}\}$ wrt s)
- > however, calculating these terms does not pose a great computational challenge



▷ an issue with this approximation is that it cannot be used for terms for which $\partial A(\hat{z}_m)/\partial z_m = 0 \Rightarrow$ it cannot be used to calculate the log-likelihood (e.g., to perform LRTs)



 \bullet Let $S(\pmb{\theta})$ true score vector; $\widetilde{S}(\pmb{\theta})$ Laplace-based score vector

$$n^{-1}S(\boldsymbol{\theta}) = n^{-1}\widetilde{S}(\boldsymbol{\theta}) + O\left\{\min(p_i)^{-2}\right\}$$

• Let $\boldsymbol{\theta}_0$ true parameter vector; $\widehat{\boldsymbol{\theta}}$ Laplace-based MLE

$$n^{-1}S(\widehat{\boldsymbol{\theta}}) = n^{-1}S(\boldsymbol{\theta}_0) + n^{-1}H(\boldsymbol{\theta}_0)(\widehat{\boldsymbol{\theta}} - \boldsymbol{\theta}_0) + O_p(1) \Rightarrow$$
$$(\widehat{\boldsymbol{\theta}} - \boldsymbol{\theta}_0) = O_p\left[\max\left\{n^{-1/2}, \min(p_i)^{-2}\right\}\right]$$

• $\widehat{\boldsymbol{\theta}}$ consistent as both $n, p_i \to \infty$



- Results from the similar framework of joint modelling of longitudinal and time-to-event data
 - \triangleright Gauss-Hermite requires creating a design matrix of dimensions $N \times h^q$ (N: total sample size; h: quadrature points; q: dimension of integration)
 - \triangleright for a data set h=3, q=8 we need 58531×6561 design matrix
 - \triangleright One EM iteration
 - * Gauss-Hermite: > 15min
 - * Fully Exponential Laplace Approximation: 12sec



- What has been done
 - \triangleright theory almost finalized
 - ▷ preliminary R programs written
- \bullet What needs to be done
 - ▷ finalize programs
 - \triangleright simulation studies

Thank you for your attention!